

# Advances in statistical methods to handle large data sets for GWAS in crop breeding

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June 4, 2018

## 1 Introduction

2 Mathew et al. (2018b) Quantitative trait loci (QTL) analysis is a well known statistical  
3 exercise in biological research to identify genetic loci associated with a quantitative  
4 trait/phenotype of interest. QTL mapping studies utilize molecular markers to locate the  
5 genomic regions that affect the phenotype. Two of the most commonly used QTL mapping  
6 approaches are linkage analysis (LA)(also known as family-based linkage mapping or  
7 QTL mapping) and association mapping (linkage disequilibrium (LD) mapping). Linkage  
8 mapping considers the linkage disequilibrium that exists within families in order to map  
9 the region, whereas the association mapping needs markers that are in LD with a potential  
10 QTL across the entire population. Association mapping is based on the assumption that  
11 the alleles which influence on the trait are inherited from a single common ancestor in the  
12 past. Table 1 summarizes a comparison between association and linkage mapping.

13 Even though, association and linkage mapping are viewed as fundamentally different  
14 approaches, both methods try to make use of the recombination events. Over the decades

**Table 1** A comparison between association and linkage mapping

Property of the mapping method	Linkage mapping	Association Mapping
Mapping populations	Close relatives	Unrelated or related individuals
Marker density	Moderate marker density	High marker density
Mapping resolution	Long (< 5 MB)	Short(< 1 Mb)
Susceptible to population stratification	No	Yes
Biological basis	Recent recombination	Historical recombination
Suitable phenotypes	Rare phenotypes	Common phenotypes
Parameter of interest	Recombination fraction	Statistical association
Controlled experiment	Yes	No

15 many LA studies (i.e., QTL mapping in offspring population resulting from a simple line  
 16 crossing experiments) have reported hundreds of QTLs in various plant species and only a  
 17 few identified QTLs were targeted at gene level (Patterson et al. (2006)). Recent advances  
 18 in low cost high throughput DNA sequencing technologies have helped genome-wide  
 19 association mapping (GWAS) to emerge as an alternative to linkage mapping and which  
 20 offers high mapping resolution and is more time-efficient. However, before starting, one  
 21 should make sure that in order to fully utilize all the potential available in GWAS, the  
 22 association mapping should be optimally performed in multiparental populations with  
 23 enough number of generations (to accumulate enough recombination events). In this  
 24 chapter we shortly discuss some of the main challenges for GWAS studies with large  
 25 datasets.

## 26 **Single-locus association model**

27 Despite the availability of large number of single-nucleotide polymorphisms (SNPs),  
 28 standard GWAS analysis methods consider one SNP at a time and identify the marker-  
 29 trait association using a single-locus model. Single-locus model is the simplest and most  
 30 commonly used model to identify associations between SNPs and continuous trait of  
 31 interest. However, it is already known that hidden population structure due to LD and of  
 32 sample structure (cryptic relatedness) leads to inflated test statistics and that may lead  
 33 to false positive and negative associations between marker and trait. Plenty of correction  
 34 methods have been proposed especially for single marker association testing where a  
 35 phenotypic relevance of a single putative gene position is tested at a time in isolation of

36 other putative loci (*e.g.*, Principal component analysis, (Price et al. 2006); Mixed-model  
37 approach, (Kang et al. 2008a, Müller et al. 2011, Yu et al. 2006); Structured association,  
38 (Pritchard et al. 2000)), for a review see (Sillanpää 2011). Mixed model method including  
39 a random polygenic term, which describes relationships between individuals, is performing  
40 well and is most widely used method in plant, animal and human datasets. It is now  
41 generally accepted that it can correct confounded (spurious) associations due to both:  
42 close relatives and population structure in the dataset. However, its general drawback is  
43 that it may loose statistical power (by over-correcting the structure) or it may lead to  
44 wrong findings if the candidate SNP is included in the calculation of genomic relationship  
45 matrix (Würschum and Kraft (2015)). A single-locus mixed model with the polygenic  
46 random effect can be expressed as:

$$\mathbf{Y} = \mathbf{X}\boldsymbol{\beta} + \mathbf{W}\mathbf{v} + \mathbf{Z}\mathbf{g} + \boldsymbol{\epsilon}. \quad (1)$$

47 Here  $\mathbf{Y} = \{Y_i\}_{i=1}^n$  is a vector of phenotype values for  $n$  lines and  $\boldsymbol{\beta}$  is the vector of fixed  
48 effects with known incidence matrix  $\mathbf{X}$ , whereas,  $\mathbf{W}$  is the incidence matrix for the marker  
49 being tested for the association. Moreover  $\mathbf{g}$  is an  $n \times 1$  vector of polygenic effects with  
50 the incidence matrix  $\mathbf{Z}$  and  $\mathbf{g} \sim N(\mathbf{0}, \mathbf{K}\sigma_a^2)$ . Here,  $\mathbf{K}$  defines the covariance structure  
51 that describes the relatedness among individuals and can be calculated either from the  
52 marker information or with the pedigree. Additionally,  $\boldsymbol{\epsilon}$  corresponds to the vector of  
53 residual, following a normal distribution as  $\boldsymbol{\epsilon} \sim N(0, \mathbf{I}\sigma_e^2)$ . In single-locus model, the  
54 marker association is tested for one marker at a time with the null hypothesis that is  
55  $v = 0$  against the alternative hypothesis is that  $v \neq 0$ . Different mixed-model methods  
56 are mainly differing in how they implement the required speed up of the computation  
57 (refitting the model and estimating the polygenic variance separately for each candidate  
58 SNP) which makes it applicable for large genome-wide datasets.

59 Some of the interesting packages based on single-locus association model are: PLINK  
60 (Purcell et al. 2007); TASSEL (Bradbury et al. 2007); CMLM (Zhang et al. 2010); ECMLM  
61 (Li et al. 2014); FaST-LMM (Lippert et al. 2011); EMMA (Kang et al. 2008b); GEMMA  
62 (Zhou and Stephens (2012)); GRAMMAR-Gamma (Svishcheva et al. 2012); rrBLUP

63 (Endelman (2011)).

## 64 **Multilocus association model**

65 Single-locus model test a single SNP at a time and known to have some drawbacks. Firstly,  
66 it is hard to locate the region contain the QTL in to a small region because a number of  
67 SNPs can be in LD with the QTL, in this case the significant SNP can span a wide range on  
68 the chromosome (Pryce et al. (2010)). Secondly, effect of a single SNP may be quite small,  
69 but may have strong joint effects and by considering all SNPs simultaneously will improve  
70 the power to detect their joint activity. One solution to these problems is to jointly fit  
71 all SNPs using a multilocus association model. Another interesting benefit of multilocus  
72 model is their capability of automatically correcting/controlling the confounding due  
73 to population structure/relatedness without having polygenic term in the model (for  
74 example, Pikkukookana and Sillanpää (2009), Würschum and Kraft (2015), Kärkkäinen  
75 and Sillanpää (2012)). Moreover, they are also relatively robust for model misspecification.  
76 The basic multilocus association model can be defined as:

$$\mathbf{Y} = \mu + \sum_{j=1}^m \mathbf{M}_{.j} \mathbf{b}_j + \boldsymbol{\epsilon}. \quad (2)$$

77 Here  $\mathbf{Y} = \{Y_i\}_{i=1}^n$  is a vector of phenotype values of  $n$  lines,  $m$  is the total number of  
78 markers,  $\mathbf{M}_{ij}$  (note that  $\mathbf{X}$  is commonly used notation for genotypes, however to avoid the  
79 confusion with the fixed effect in linear mixed model we use the notation  $\mathbf{M}$  here) is the  
80 genotypic value of individual  $i$  at marker  $j$  coded as 0, 1, 2 for the genotype  $AA$ ,  $Aa$ ,  $aa$   
81 respectively,  $\mathbf{b}_j$  (note that  $\boldsymbol{\beta}$  is the commonly used notation for marker effect and to avoid  
82 the confusion with the fixed effect term in Eq. 1 we use the notation  $\mathbf{b}$ ) is the random  
83 marker effect associated with marker  $j$ , and  $\boldsymbol{\epsilon}$  corresponds to the residual, following a  
84 normal distribution as  $\boldsymbol{\epsilon} \sim N(0, \mathbf{I}\sigma_e^2)$ . With multilocus model regression methods are  
85 generally used to estimate marker effects ( $\mathbf{b}$ ).

86 Some of the interesting packages based on multilocus association model are: MLMM  
87 (Segura et al. 2012); FASTmrEMMA (Wen et al. 2017); MRMLM (Wang et al. 2016;

88 E-Bayes (Xu 2007).

## 89 **High dimensional data space in GWAS**

90 Nowadays GWAS studies involves thousands of SNPs owes much to the recent advances  
91 in genotyping technologies. This availability of high-throughput genomic data leads  
92 to the ‘large  $p$ , small  $n$ ’ (here  $p$  corresponds to the number of marker effects and  $n$  is  
93 for the number of samples) problem in GWAS. The so called ‘large  $p$ , small  $n$ ’ occur  
94 when the number of parameters to be estimated (marker effects  $\mathbf{b}$  in model 2) is much  
95 larger than the samples (phenotype values) and biologists have to deal with large data  
96 space. Additionally, strong LD among SNPs poses additional challenges in GWAS studies.  
97 Multilocus models are more prone to ‘large  $p$ , small  $n$ ’ problem because joint analysis of  
98 all SNPs together is computationally challenging, whereas the one-dimensional genome  
99 scan by testing a single SNP at a time can handle a large number of markers without  
100 much problems. Two of the most commonly used methods to deal with large data space  
101 are regularization and variable selection methods.

## 102 **Variable selection and shrinkage/regularization**

103 Identifying the relevant predictive variables is a fundamental problem in statistical learning.  
104 Most standard regression methods may fail in such cases where the number of markers is  
105 much larger than the sample size. Variable selection and shrinkage (regularization) methods  
106 are commonly applied to such problems to select the best subset of predictors. Stepwise  
107 procedures (forward selection and backward elimination) are commonly used for variable  
108 selection in ‘large  $p$ , small  $n$ ’ problems. Backward-forward variable selection methods  
109 are only applicable with a couple of tens of markers and become quickly impractical  
110 as the number of predictors increases. The single-locus model can also be extended to  
111 the multilocus framework by applying stepwise (forward/backward) regression approach  
112 proposed by Segura et al. (2012). Stepwise regression is an iterative procedure, where in  
113 each step, a SNP is added to the model as a cofactor based on predefined criteria. Then  
114 the  $p$ -values for all added cofactors are re-estimated together with the variance components.

115 The process of adding significant SNPs to the model is repeated until the benefit of adding  
116 new terms to the model comes sufficiently close to zero. Shrinkage/regularization methods  
117 are another class of estimation approach used to solve the ‘large  $p$ , small  $n$ ’ problems.  
118 Regularization methods attempt to estimate all the genetic effects, while the effects of  
119 irrelevant covariates (spurious effects) are automatically shrunken toward zero.

## 120 **Frequentist regularization approaches**

121 Ridge regression and LASSO (least absolute shrinkage and selection operator)(Tibshirani  
122 (1996)) are prominent shrinkage methods in the classical framework, and fall under the  
123 umbrella of penalized likelihood regression models, with the penalty being imposed on  
124 the L2-norm for the former and on the L1 -norm for the latter (see Schmidt (2005) for  
125 more details). In frequentist framework, LASSO and its extensions adaptive LASSO (Zou  
126 (2006)), elastic net (Zou and Hastie (2005)) and adaptive elastic net (Zou and Zhang  
127 (2009)) have been widely used for association mapping or genomic selection studies (Chen  
128 and Chen (2008); Wang et al. (2011); Waldmann et al. (2013); Sokolov et al. (2016)). Fan  
129 and Li (2001) showed that the LASSO shrinkage produces biased estimates for the large  
130 coefficients, and Zou (2006) proposed an extension of LASSO called adaptive LASSO in  
131 order to alleviate this bias. Another limitation of LASSO approach is that when there exist  
132 high correlations among predictors LASSO will arbitrarily choose one and drop the other  
133 predictors. To remedy this problem, elastic net (ENET) was proposed as an extension  
134 of LASSO and ENET is robust to high correlations among predictors. Later, Zou and  
135 Zhang (2009) proposed adaptive elastic net as a combination of the adaptive LASSO and  
136 the elastic net to deal with the collinearity problem and improved performance with high-  
137 dimensional data. See Ogutu et al. (2012) and Li and Sillanpää (2012) for a comprehensive  
138 review about the frequentist regularization procedures in association mapping and genomic  
139 selection. It has been long argued that the classical shrinkage methods (LASSO and  
140 its extensions) cannot identify a number of non-zero effects exceeding the sample size.  
141 This is a major shortcoming when dealing with genome-wide dense sets of markers and  
142 Bayesian formulations of the regularization methods can overcome this with the help of

143 prior distributions.

## 144 **Bayesian regularization approaches**

145 For the Bayesian inference with model 2 (Eq. 2), prior distributions must be specified for  
146 the unknown parameter such as  $\mathbf{b}_j$  and  $\sigma_e^2$ . In the Bayesian framework, regularization is  
147 achieved by imposing specific prior distribution on the random marker effects and the  
148 priors shrink unimportant marker effects toward zero. Normal distribution with a common  
149 variance is the simplest prior one can be assume for SNP effects and this is equivalent to  
150 the ridge regression BLUP. One disadvantage of using normal distribution for SNP effects  
151 is that finally large number of SNP effects will have non-zero values. As a solution to  
152 this problem some heavy-tailed distribution, like t-distribution, can be used as a prior  
153 distribution for the SNP effects (BayesA; Meuwissen et al. (2001)). Another commonly  
154 used heavy-tailed shrinkage distribution for SNP effects is Laplace (double exponential)  
155 distribution, which is sharply peaked around zero. This is known as Bayesian LASSO:  
156 Park and Casella (2008), Li et al. (2010) and its adaptive counterpart: Extended Bayesian  
157 LASSO; Mutshinda and Sillanpää (2010)). Many other variants exist including Meuwissen  
158 et al. 2001 (BayesB) and Habier et al. 2011 (BayesC and BayesC $\pi$ ). These multilocus  
159 models can be used both for association mapping and for genomic prediction.

160 The parameter estimation in most of the Bayesian hierarchical shrinkage methods is  
161 based on Markov chain Monte Carlo (MCMC) sampling which may not be optimal for  
162 high dimensional data due to their computational complexity. Deterministic methods  
163 such as *maximum a posteriori* (MAP) estimation can be a used as an fast alternative to  
164 sampling based algorithms. However, MAP estimation methods can produce good point  
165 estimates but their accuracy estimates are usually badly underestimated (*i.e.*, the estimated  
166 posterior uncertainty around the point estimate is much too narrow). MAP estimation is  
167 mainly based on numerical optimization (Gelman et al. (2014)) or different variants of  
168 expectation maximization (EM)(Dempster et al. (1977)) algorithm. Variational Bayes (VB)  
169 estimation (Jaakkola and Jordan (2000)) offers another class of MAP estimation technique  
170 in multilocus models (Li and Sillanpää (2012)) but also their uncertainty measures

171 are too narrow. Variational Bayes can be considered as the extension of traditional  
172 expectation-maximization (EM) algorithm and is computationally less intensive than  
173 MCMC counterparts for the shrinkage estimation. Many MAP implementations exist for  
174 large data sets (for example, Sun et al. (2010), Zhang and Xu (2005), Huang et al. (2015),  
175 Mutshinda and Sillanpää (2012), Li and Sillanpää (2012)).

## 176 **Significance threshold for association**

177 Even though GWAS studies have great potential to pinpoint the single nucleotide polymor-  
178 phisms underlying quantitative traits, false discoveries are a major concern in associations  
179 studies. In a single-locus model-based GWAS study, one is typically screening through  
180 thousands of markers by testing association one at a time which may lead to many false  
181 positive findings. One important question is which significance level ( $\alpha$ ) should be chosen  
182 in order to reduce the number of false alarms due to multiple testing (i.e., high number  
183 of tests performed). Bonferroni correction is one of the most commonly used method to  
184 correct for multiple testing in GWAS studies. Bonferroni adjustment treats all markers as  
185 independent, even though the markers are likely to be in LD with each other. Thus the  
186 Bonferroni adjustment may be too conservative for extremely large number of markers.  
187 An alternative to Bonferroni correction, FDR (false discovery rate), which is designed to  
188 capture the portion of false positives to the number of total positive test results is also  
189 widely applied in GWAS studies (Devlin et al. (2003)). The quantile-quantile (Q-Q) plot  
190 (which is a graphical representation of the proportion of significant SNPs compared to the  
191 expected number of significant SNPs based on  $p$ -values) is also used in GWAS studies  
192 based on single-locus model to monitor the number of false positives.

193 Bonferroni correction and Q-Q plot are suitable when the loci are independent. Thus,  
194 it is hard to apply these methods for the multilocus association models because such  
195 models search loci jointly and their combinations can be correlated. Permutation test  
196 proposed by Churchill and Doerge (1994) is commonly applied for choosing the significance  
197 level for association in both single and multilocus association analysis (Xu (2003)). As an  
198 interesting alternatives credible interval approach (Li et al. (2010)) and Wald-statistic

199 (Yang and Xu (2007)) can also be used in multilocus association testing to decide which  
200 signals are true. But all these approaches are generally sensitive to the collinearity in the  
201 marker data and among these methods permutation test seems to suffers less from the  
202 correlated predictors. Another interesting approach is to use the estimates from many  
203 MCMC chains with different starting values, and consider only the SNPs (stable signals)  
204 that appear in all MCMC iterations as the significant ones (Mathew et al. (2018a); Wei  
205 et al. (2014)). For more alternatives, see Chen et al. (2017).

## 206 **Dimensionality reduction methods**

207 The cost of high-throughput genotyping/phenotyping is no longer a major hurdle for  
208 GWAS studies and the biologists have entered the era of Big Data. Variable selection  
209 and shrinkage methods are mainly designed to moderate the number of predictors to  
210 hundreds or thousands. However with Big Data (ultra-high  $p$  small  $n$ ), these methods may  
211 be computationally infeasible and statistically inaccurate. Another problem associated  
212 with high dimensional data is that many unimportant predictors can be highly correlated  
213 with the response variable and variable selection alone might be difficult in such cases.  
214 While making statistical inference with Big data (high dimensional data space) one can  
215 use dimensionality reduction approach to reduce the number of predictors ( $p$ ) close to  
216 the sample size before applying variable selection/regularization methods. Collinearity,  
217 which is a condition where some of the predictors are highly correlated due to LD with  
218 each other is a major problem with the high dimensional data space. In such cases LD  
219 pruning can be applied to remove the highly correlated SNPs and preselect a subset of  
220 SNPs which are uncorrelated with each other. Then the selected subset of SNPs can  
221 be further analyzed using a multilocus association model. The PLINK software offers  
222 features for SNP pruning based on LD. SNP tagging (Lin and Altman (2004); Meng et al.  
223 (2003)) and SNP binning (Xu (2013)) based on haplotypes are another useful approaches  
224 to reduce the dimensionality (by selecting an informative sets of SNPs) in GWAS studies  
225 involving millions of SNPs. Sure Independence Screening (SIS) (Fan and Lv (2008)) is  
226 an efficient method to reduce the dimensionality of high dimensional data space from

227 ultra-high to a relatively large scale. SIS can preselect the most important predictors based  
228 on their marginal correlation with the response variable. Recent studies (Kärkkäinen  
229 et al. (2015); Mathew et al. (2018a)) showed that SIS can be applied to preselect the  
230 important predictors for multilocus association in very high dimensional cases. SIS is  
231 based on univariate screening step and one of the drawback of SIS is that the important  
232 predictors that are marginally nearly uncorrelated with the response variable could be  
233 missed out because of this univariate screening approach. In order to over come this  
234 drawback Fan and Lv (2008) also proposed an iterative procedure called iterative sure  
235 independence screening (ISIS). The ISIS procedure iterates the SIS procedure conditional  
236 on the previously selected predictors, thus the procedure can capture the important  
237 predictors that are marginally nearly uncorrelated with the response variable.

## 238 **Conclusion**

239 High throughput genotyping technologies are capable of generating enormous set of high  
240 density SNP markers with low cost and that enables whole-genome association mapping  
241 in natural/breeding populations. Multilocus association mapping approaches, known  
242 to have some advantages over conventional QTL-mapping, may importantly have more  
243 power to detect QTLs and to control the number of false positives. However, multilocus  
244 association analysis involving high-density markers need to apply variable selection or  
245 shrinkage approaches in order to identify the best subset of relevant predictor variables.  
246 Even though most of the variable selection/shrinkage approaches presented in plant or  
247 animal genetics context are primary designed for genomic prediction purposes, they can  
248 also be applied for gene mapping. When such methods are applied for association analysis  
249 based on multilocus association models one need to perform additional statistical tests for  
250 association between the SNPs and the trait of interest to fully control false positives.

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